AMS 7: Statistical Methods For the Biological, Environmental and Health Sciences

1: Introduction and Descriptive Methods

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Outline

- Introduction: populations and samples; parameters and statistics (estimates)
- **Data types**: qualitative and quantitative variables; nominal and ordinal; discrete and continuous; interval and ratio; dichotomous
- Descriptive methods for a single variable
 - Graphical: histograms, bar charts
 - Numerical: measures of center (mean, median, mode) and spread (standard deviation, variance)
- Using the normal distribution descriptively

1.1 Introduction

<u>Statistics</u> is the **study of uncertainty**: how to **measure it**, and what to **do about it**.

Uncertainty is a state of incomplete or imperfect information about something of interest to you, for example

the percentage θ of the deer who live on the UCSC campus as of 31 December 2006 who have chronic wasting disease. 31 May 19

I notice that I don't know the value of θ exactly; I have the impression that θ is rather small, since the deer on campus seem relatively healthy, but I have substantial uncertainty about its precise value.

I can reduce my uncertainty by gathering data on the disease status of campus deer; how should this data-gathering be done?

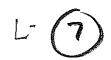
The set

 $\mathcal{P} = \{ \text{the deer who live on the UCSC campus} \text{ as of } \frac{\mathbf{31 \cdot December}}{\mathbf{31 \cdot Pq}} \frac{\mathbf{2006}}{\mathbf{3000}} \}$

is an example of a **population**: a collection of **subjects** or **elements** (in this case, deer) of interest to me.

There is an aspect of each of these population subjects that I'm curious about: if I encountered one of these deer, the question I would ask is "Does this deer have chronic wasting disease or not?"

Things that can be **measured** on population subjects are called **variables**; in this case the variable of interest takes on only two values, {**yes**, **no**} (such variables are called **dichotomous** or **binary**).



Populations and Samples; Parameters and Statistics (Estimates)

We'll see soon that a **handy** way to work with dichotomous variables is to assign 1 and 0 to their two possible values (hence the term **binary**); for example, with the variable (**chronic wasting disease or not**) the coding (1 = yes, 0 = no) is particularly useful.

A numerical summary of a population is called a parameter; θ is an example of one possible parameter of interest about the population \mathcal{P} above (others might include the average weight of the deer who are more than three months old).

If I had enough time and money (and a way of ensuring that I could find all the deer and mark them uniquely, so that I didn't double-count any individual), in principle I could perform a complete census of the entire population, obtaining the disease status for each individual, and at the end of this census I would no longer have any uncertainty about the parameter θ .

In practice people **rarely** have enough time and money to perform a **complete census** of a population \mathcal{P} ; instead it's natural to choose a **subset** \mathcal{S} of \mathcal{P} and evaluate the variable(s) of interest **only on the population subjects in the subset**.

Such a subset is called a <u>sample</u> from the population \mathcal{P} — if the sample is chosen well, it seems like a good idea to use the data in the sample to make an <u>estimate</u> of (an educated guess at) the population parameter θ of interest.

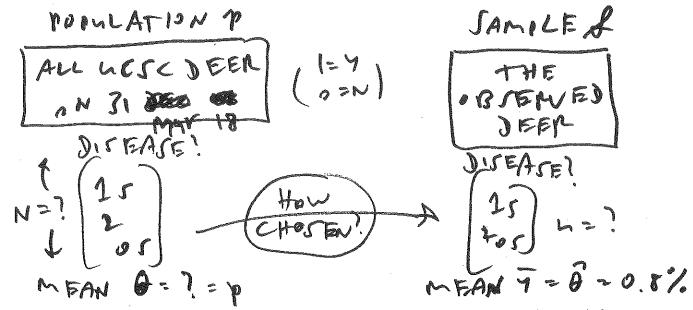
An estimate $\hat{\theta}$ of a population parameter θ is also sometimes called a **statistic**.



Populations and Samples

Let's let N stand for the **number of subjects** in the **population** \mathcal{P} , and n denote the **number of subjects** in the sample \mathcal{S} .

Then both the **population** and the **sample** can be thought of as **data sets**, which can be further visualized as **rectangular tables*** with **one row** for each **subject** and **one column** for each **variable**, as in the following **diagram**:



In this class we'll be looking a lot at diagrams like this one: such diagrams are the basis of both **probability models** and **statistical models**, both of which are crucial to the process of **quantifying uncertainty**.

To fill out a diagram like this I need to specify the following **ingredients**:

In the box above the population data set P I describe the subjects in P by saying to myself "There's one row in the population for each ____" and filling in the blank; for example, here there's one row for each deer living on the Santa Cruz campus on 31 December 2006.

*Math note: the official name for such a rectangular table is a matrix.



Random Sampling

- Above each **column** in the population data set I write the name of the **variable** summarized by that column (in this part of the class we'll typically work with only **one variable** at a time); here the variable of interest is the answer to the question (chronic wasting disease?).
- I identify the **number** N of subjects in the population if I know it (here I'm **not sure** how many deer there are on the UCSC campus in December 2006, so I just put a **question mark**).
- Then I do the same three things for the **sample** data set: in the box above S I describe the **subjects** in the **sample** (here I might just say "the observed deer"); above each column in the sample data set I write the name of the variable summarized by that column (this will be the **same** as in the population); and I identify the number n of subjects in the **sample** if I know it (here we haven't yet talked about how large the sample should be, so again I just put a question mark).

There's one crucial thing about this concept of **using the** sample data to estimate a parameter of interest in the population: I said above that this is a good idea "if the sample is chosen well," and we need to figure out what this means.

Evidently, if the sample is to serve as a **good stand-in** for the rest of the population, the **basic principle** we want to follow is to try to make the **sample** and the **unsample** (the part of the population not chosen in the sample) **as similar** as possible in all relevant ways.

The simplest way to achieve this goal turns out to be to draw the sample at random from the population (so that all subsets have an equal chance of being chosen).

SRS and IID Sampling

So the last (and perhaps **most crucial**) step in filling out the **diagram** above is as follows:

• Finally, in the circle above the arrow from the population to the sample I describe the sampling method (in this case, random).

To literally take a random sample of size n of deer from \mathcal{P} , you'd have to

- (a) make a **list** of all the **population subjects** (deer), with **unique identifying tags**;
- (b) choose n of these tags at random (using, for example, pseudo-random numbers generated by computer) without replacement (the sampling method at random without replacement is called simple random sampling (SRS), as opposed to at random with replacement, which is called independent identically distributed (IID*) sampling); and
- (c) measure the **variables** of interest on the sampled deer by finding the ones with the **chosen tags**.

In practice people would often instead use a **simpler** method that's not literally SRS (for example, if the deer were well distributed spatially, you could **partition** the UCSC campus into *n* non-overlapping and exhaustive spatial subsets and have *n* people each get data on the **first deer they encounter** in their subset on a given day) and then argue that their simpler method was **like** what you would get with SRS.

^{*}Textbook note: IID is a term not mentioned in Triola & Triola.



1.2 Data Types

It's useful to have a **classification** of the various **types of data** that variables can keep track of (because some
methods of analysis are definitely **not appropriate** for some
data types).

you're studying may take on only two values (brown, blue) that have no unique place on the number line (earlier we called such variables dichotomous or binary); similarly, hair color might take on four values (predominately brown, black, red or white).

Variables like this are said to occur on a <u>nominal</u> scale of measurement (so dichotomous variables with values like {yes, no} are **special cases** of nominal variables).

Example 2: Success in running a maze might be recorded

1 (very slow), 2 (slow), 3 (moderate), 4 (fast), 5 (very fast)

There are still **no unique places on the number line** for such values, but (unlike example 1) there's a **natural ordering** to these values.

Variables like this are said to occur on an ordinal scale.

Some other names for **nominal** and **ordinal** variables are **qualitative** and **categorical**.

Data Types (continued)

Example 3: Size of a plant. Two measures of the size of a plant (which, in turn, is a measure of its competitiveness) would include its height (in centimeters (cm)) and the number of leaves it has.

Unlike the situations in Examples 1 and 2, the values taken on by these variables do have unique places on the number line, and in fact there are two important characteristics of the numerical values of these variables:

- (a) there is a **constant size interval** between any adjacent units on the measurement scale, so that the concept of $\bf 1$ unit means the same thing anywhere on the scale (for example, plants A, B, C and D are (respectively) 14, 15, 62, and 63 cm high; the **amount** by which B is **taller** than A is the same as the **amount** by which D is **taller** than C); and
- (b) there is a **true zero** on the measurement scale with a **direct physical meaning** this allows us to make meaningful statements about **ratios** (for example, plant C is $\frac{62\text{cm}}{15\text{cm}} \doteq 4.1$ times taller than plant B).

Variables like this are said to occur on a ratio scale.

Example 4: Growing temperature at which a plant produces the most buds. Temperature (measured either in $^{\circ}C$ or $^{\circ}F$) does have a **constant size interval** but **lacks a true zero**, so (contrary to statements you see in the newspaper or on TV) when it's $80^{\circ}F$ outside you can't correctly say that it's **twice as hot** as when it's $40^{\circ}F$.

Variables like this are said to occur on an interval scale.

Data Types (continued)

Some other names for **ratio** and **interval** variables are **quantitative** and **numerical**.

One last distinction: plant height and number of leaves are different in that with plant height, conceptually (with finer and finer measuring instruments) there are no possible gaps between the possible values, whereas with number of leaves, distinct structural gaps exist (it doesn't make sense to talk about $4\frac{1}{2}$ leaves).

Quantitative variables with **gaps** between the possible values are called **discrete**; quantitative variables with **no conceptual gaps** between the possible values are called **continuous**.

Why these distinctions matter. Suppose I choose to code the age of some animals I'm observing in the following way, when storing the values of this variable in a computer:

less than 1 year old = 0, between 1 and 2 years old = 1, between 2 and 3 years old = 2, between 3 and 4 years old = 3, ...

Suppose further that I choose to **code** the hair color of these animals in the following way:

brown = 0, black = 1, red = 2, white = 3

Here's the data set I get (written, to save space, in a **transposed** fashion in relation to the convention on page 5 above: here the **rows** are the **variables** and the **columns** are the **subjects** (animals)):

						Mean
Animal Identifier	45	333	167	2	 501	243.9
Age	2	0	1	1	 3	1.7
Hair Color	0	2	3	2	 1	1.2

1.3 Descriptive Methods

As we'll soon discuss, it's sometimes both **useful** and **meaningful** to **summarize** a variable by taking its **mean** (just add 'em up and divide by how many there are); the computer has done this for us in the table above in the **final column**.

The problem is, of course, that the **mean** is **meaningful** only for the **age** variable (because it's **quantitative** [ratio, discrete]; the other two variables are **qualitative** [nominal]).

The point: The right way to analyze a variable often depends on the scale on which it's measured.

1.3.1 Graphical descriptive methods. Example:

butterfly wing lengths. Zar (1999) gives data from a sample of n=24 immature monarch butterflies, in which the variable of interest (we might call it y; T&T would call it x) is wing length (in cm):

4.4 3.6 4.1 3.3 3.5 3.8 4.5 4.3 4.3 4.0 4.1 3.6 4.0 4.0 3.8 3.8 3.9 4.2 4.2 4.1 3.7 3.9 4.0 3.9

(This is just shorthand for a data set with n=24 rows (subjects = butterflies) and 1 column (variable = wing length), written in this manner to save space.)

How might we summarize this variable in a way that would allow us to see patterns (graphical summaries) and to capture most of the information it contains in fewer than 24 numbers (numerical summaries)?

Raw Frequency Distribution

As long as the **order** in which the data values were listed above is **not relevant**, the first step would be to **sort** the data from **smallest** to **largest**:

Now we can see that there are a number of **duplicate values** (caused by **rounding** the wing length measurement to the nearest cm).

This suggests a **further summary** in which we keep track of the **values** of the variable and the **raw frequencies** (the numbers of times those values are attained):

Value	Frequency
3.3	1
3.4	0
3.5	1
3.6	2
3.7	1
3.8	3
3.9	3
4.0	4
4.1	3
4.2	2
4.3	2
4.4	1
4.5	1
Total	n = 24

This is called a <u>raw frequency distribution</u> (or <u>frequency table</u>) for the variable y (sometimes people just refer to the <u>distribution</u> of y, or ask "How is y <u>distributed?"</u>).

Raw Frequency Histogram

The **table** on the previous page is not as easy to **absorb** as it would be if we could display it **graphically**.

Since it has **two columns** or **dimensions**, it's natural to make a plot in which one dimension (**horizontal**, say) is the **values** the variable takes on and the other (**vertical**, say) is the **raw frequencies** — the resulting graph is a (raw frequency) **histogram** of the variable y:

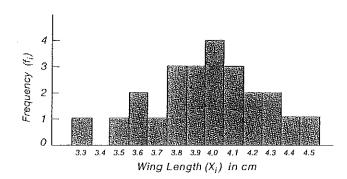


Figure 3.1 A histogram of the data in Example 3.2. The mean (3.96 cm) is the center of gravity of the histogram, and the median (3.975 cm) divides the histogram into two equal areas.

A histogram is a special case of a <u>bar graph</u>: a plot with locations identified along the horizontal axis corresponding to values a variable takes on and <u>bars</u> over those locations with <u>heights</u> give by the (raw) <u>frequencies</u> of those values.

A bar graph can be drawn as a summary of any qualitative (nominal or ordinal) variable; there is no unique place called "yes" or "red" on the number line, but you can just invent arbitrary horizontal locations and make a useful plot anyway.

Strictly speaking, what makes a **histogram** a histogram is that the variable in question is **quantitative** (so that the values do have unique locations on the number line) — histograms can be made for either **discrete** or **continuous** variables.

More Graphical Examples

	Number of
Nest site	nests observed
A. Vines	56
B. Building eaves	60
C. Low tree branches	46
D. Tree and building cavition	es 49

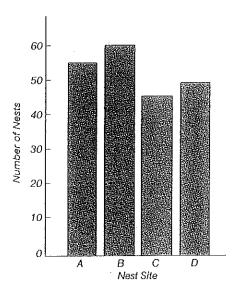


Figure 1.1 A bar graph of the sparrow nest data of Example 1.1. An example of a bar graph for nominal data.

EXAMPLE 1.2 Numbers of sunfish, tabulated according to amount of black pigmentation. A frequency table of ordinal data.

Pigmentation class	Amount of pigmentation	Number of fish
0	No black pigmentation	13
1	Faintly speckled	68
2	Moderately speckled	44
3	Heavily speckled	21
4	Solid black pigmentation	8

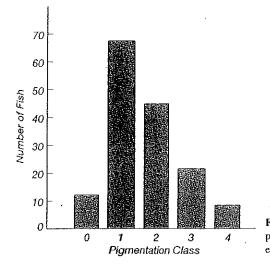


Figure 1.3 A bar graph of the sunfish pigmentation data of Example 1.2. An example of a bar graph for ordinal data.



Graphical Examples (continued)

EXAMPLE 1.3 Frequency of occurrence of various litter sizes in foxes. A frequency table of discrete, ratio-scale data.

Litter size	Frequency
3	10
4	27
5	22
6	4
7	i

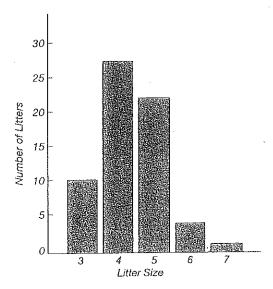


Figure 1.4 A bar graph of the fox litter data of Example 1.3. An example of a bar graph for discrete, ratio-scale data.

EXAMPLE 1.4a Number of aphids observed per clover plant. A frequency table of discrete, ratio-scale data.

Number of aphids on a plant	Number of plants observed	Number of aphids on a plant	Number of plants observed
Q	3	20	17
Ì	Ī	-21	18
2	1	22	23
3	1	23	17
4	2	24	19
5	3	25	18
6	5	26	19
7	7	27	21
. 8	8	28	18
9	11	29	13
10	10	30	10
11	11	31	14
12 .	13	32	9
13	12	33	10
14	16	34	8
15 .	13	35	5
16-	14	36	4
17	16	37	1
18	15	38	2
19	14	39	1
		40	0
		41	1

Graphical Examples (continued)

EXAMPLE 1.4b Number of aphids observed per clover plant. A frequency table grouping the discrete, ratio-scale data of Example 1.4a.

Number of aphids on a plant	Number of plants observed
0–3	6
4–7	17
8-11	40
1215	54
16-19	59
20-23	75
24–27	77
28-31	55
32-35	32
36-39	8
40-43	1

Total number of observations = 424

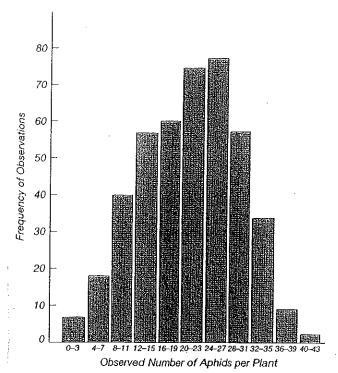


Figure 1.5 A bar graph of the aphid data of Example 1.4b. An example of a bar graph for grouped discrete, ratio-scale data.

Graphical Examples (continued)

		Frequency	Cumulativ	Cumulative frequency		
	Phosphorus mg/g of leaf)	(i.e., number of determinations)	Starting with low values	Starting with high values	-	
	8.15-8.25	2	2	130	,	
	8.25-8.35	6	8	128		
	8.35-8.45	8	16	122		
	8.45-8.55	. 11	27	114		
	8.55-8.65	17	44	103		
	8.65-8.75	17	61	86		
	8.75-8.85	. 24	85	69		
	8.85-8.95	18	103	45		
•	8.95-9.05	13	116	27		
	9.05-9.15	10	126	14		
	9.15-9.25	4	130	4		

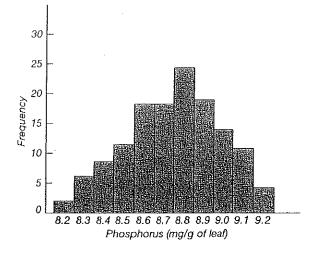


Figure 1.6a A histogram of the leaf phosphorus data of Example 1.5. An example of a histogram for continuous data.

1.3.2 Numerical Descriptive Methods

In addition to **summarizing** a variable **graphically** to look for **patterns**, it's also useful to **summarize** it **numerically**, to capture **most of the information it contains** in **fewer than** *n* **numbers**.

People have found two main types of numerical summary useful: measures of center (or central tendency), and measures of spread (or dispersion, or variability).

Measures of center. The three most useful are the mean, the median (and other quantiles, or percentiles), and the mode.

The <u>mean</u> is, I'm sure, an old familiar object: with a variable like **butterfly wing length** above (n = 24)

$$4.4$$
, 3.6 , 4.1 , ..., 4.0 , 3.9
 y_1 , y_2 , y_3 , ..., y_{23} , y_{24}
 y_1 , y_2 , y_3 , ..., y_{n-1} , y_n

to get the **mean** (for a variable y, let's call the mean \bar{y}) I just **add up** all the data values and **divide** by **how many** there are:

$$\bar{y} = \frac{4.4 + 3.6 + 4.1 + \dots + 4.0 + 3.9}{24} = \frac{95}{24} = 3.96 \,\text{cm}.$$
 (1)

Symbolically, using the idea of **summation notation**, this can be written more **succinctly** as

$$\bar{y} = \frac{y_1 + y_2 + y_3 + \dots + y_{n-1} + y_n}{n}$$

$$= \frac{1}{n} (y_1 + y_2 + \dots + y_n) = \frac{1}{n} \sum_{i=1}^{n} y_i.$$
 (2)

Measures of Center

It turns out that the **mean** has the **graphical interpretation** of the **center of gravity** of the data: if you visualize the **histogram** of the variable y as made of **bricks** that are sitting on a **number line** made of **plywood**, which in turn is put **on top of a saw-horse**, the mean is **the place where the histogram would exactly balance**.

The <u>median</u> \tilde{y} of a column of numbers (y_1, \ldots, y_n) is defined to be the <u>middle</u> value in the list in which (y_1, \ldots, y_n) has been <u>sorted</u> from smallest to largest — if n is an <u>odd</u> number this is <u>uniquely defined</u>; if n is <u>even</u> there is <u>no</u> single <u>middle</u> number and people define the median to be the <u>mean</u> of the two middle values.

By examining the sorted list of the butterfly wing lengths on page 12 above (n=24), you can see that the **middle** two numbers are both 4.0, so the **median** \tilde{y} of that variable is 4.0cm (not very far from the mean in this case).

Since histograms graphically display the frequency distribution of a variable, the graphical interpretation of the median must be that it's the place where half of the data is to the left of that place and half to the right in the histogram.

The **median** is a **special case** of the general idea of finding places in the distribution where a particular **percentage** of the data is **to the left** of that place — these are called **quantiles** or **percentiles**.

By definition the **median** is the **50th percentile**, but it's also useful sometimes to look at **other percentiles** — for example, the **25th percentile** (also called the **first quartile**) is the place where $\frac{1}{4}$ of the data is to the **left** of that place, and similarly the **75th percentile** (the **third quartile**) is the place where $\frac{1}{4}$ of the data is to the **right** of that place.

Mean, Median, Mode

The <u>mode</u> is defined to be the <u>point of highest frequency</u> in the distribution of a variable, so by definition its <u>graphical</u> interpretation is just the <u>highest point in the histogram</u> (Note: there are <u>many different possible histograms</u> for the same set of data [as a function of <u>how wide you</u> choose the bars to be and <u>where you decide to start the first bar</u>], and the mode is <u>sensitive</u> to these choices).

Some distributions have more than one mode — these are called multimodal (a common special case is two modes, which defines a bimodal distribution), and distributions with only one mode are therefore unimodal; usually multimodality means that there are two or more subgroups in the sample that should perhaps be studied separately.

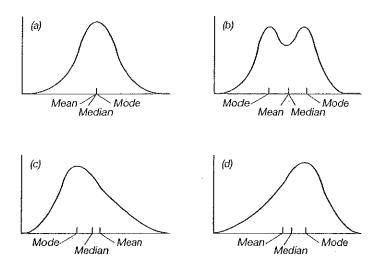


Figure 3.2 Frequency distributions showing measures of central tendency. Values of the variable are along the abscissa (horizontal axis), and the frequencies are along the ordinate (vertical axis). Distributions (a) and (b) are symmetrical, (c) is positively skewed, and (d) is negatively skewed. Distributions (a), (c), and (d) are unimodal, and distribution (b) is bimodal. In a unimodal asymmetric distribution, the median lies about one-third the distance between the mean and the mode.*

Distributions that are composed of mirror images to the left and right of a central folding point are called symmetric; combining two of these terms, a fairly common distributional shape is symmetric unimodal (like Figure (a) above) — for such distributions the mean, median and mode all coincide with the point of symmetry, which makes choosing a measure of center for them easy.

Outliers; Sensitivity Analysis

	ecies A (mo)	Species B X_i (mo)	
	34	34	
•	36	36	
	37	37	
	39	39	
	40	40	
	41 .	41	
	42	42	
	43	43	
	79	44	
	·	45	
n = 9		n = 10	
$\mathcal{M}=X_{(n+1)}$	$Y_{/2} = X_{(9+1)/2}$	$\mathcal{M} = X_{(n+1)/2} = X_{(10+1)/2}$	
$= X_5 = 0$	40 mo	$= X_{5.5} = 40.5 \text{ mo}$	
$\bar{X} = 43.4 \text{ n}$	าด	$\vec{X} = 40.1 \text{ mo}$	

The parts of a distribution to the **left** and **right** of the **center** are called the left and right <u>tails</u> of the distribution, respectively.

Notice in the data set for species A above that the largest observation in the right tail (a life expectancy of 79 months (mo)) is much larger than the others — data values in either tail that are far from the bulk of the data are called outliers.

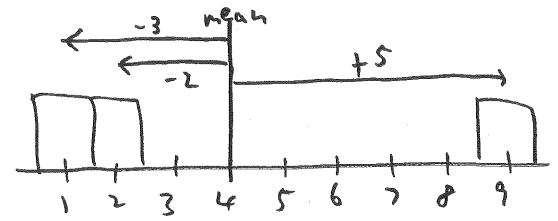
You can see that with data set A the mean has been quite strongly influenced by the outlier (the median is 40 mo, and the mean has been pulled by the outlier all the way up to 43.3 mo) — this observation (the fact that the mean is more sensitive to outliers than the median) is part of a general phenomenon.

We'll see examples later of what to do about outliers—
one of the simplest ways to address them is by means of
sensitivity analysis (repeat your main analyses with and
without the outliers and see if you get more or less the
same results— if so, great; if not, you have to think
harder about whether there's a good scientific reason to
discard the outliers).

Measures of Spread

Measures of spread. The two most useful are the variance and the standard deviation (SD).

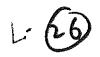
To see what's going on in **measuring** the **spread** of a list of numbers, consider the tiny **fake data set** with n=3 values $(y_1,\ldots,y_n)=(1,2,9)$, whose **mean** \bar{y} is **4** and whose **histogram** looks like this:



We might intuitively define the spread of a list of numbers to be the typical amount by which the numbers differ from their center; how should this idea be quantified?

One way might begin by calculating the <u>deviations</u> $(y_i - \bar{y})$ of each **observation** y_i from the **mean** \bar{y} :

The deviations represent the amount by which each number differs from the center (as measured by the mean); all we need to finish off the calculation is to summarize them (to get the "typical" deviation).



Variance

One way to **summarize** the deviations would be to take their **mean**, but this **doesn't work**: with the example above you get $\mathbf{0}$, and in fact by the way the mean is defined you would **always get 0** no matter what the numbers (y_1, \ldots, y_n) are (it's not hard to **prove** this **algebraically**) — the problem is **cancellation** of + and - deviations.

One way to avoid cancellation is to take the mean of the absolute values of the deviations — this is the mean absolute deviation (MAD), which here comes out $\frac{|1-4|+|2-4|+|9-4|}{3} = \frac{3+2+5}{3} \doteq 3.3$; this does seem to correspond to the **typical length** of the **arrows** in the **sketch** above.

For technical reasons having to do with calculus and theoretical statistics, the MAD is not used much; the two most frequently used measures of spread are based on the other way to avoid cancellation: taking (more or less) the mean of the squares of the deviations.

Again for **technical reasons** (which will be explained later), when the data set you're measuring the spread of is a **(random) sample** from a **population**, in calculating this "**mean**" people prefer to divide not by the number of data values n but by (n-1); the resulting quantity is called the **(sample) variance**, usually abbreviated s^2 :

$$s^{2} = \frac{(y_{1} - \bar{y})^{2} + \dots + (y_{n} - \bar{y})^{2}}{n - 1} = \frac{1}{n - 1} \sum_{i=1}^{n} (y_{i} - \bar{y})^{2}.$$
 (3)

Thinking of the little **fake data set** as a **sample**, the **(sample) variance** comes out

 $\frac{(1-4)^2+(2-4)^2+(9-4)^2}{3-1}=\frac{3^2+2^2+5^2}{2}=19$, which seems **too big** as a measure of the **typical length** of the **arrows** in the **sketch** above, and moreover **the units of the variance are wrong**: if the data values were measurements of **money** (\$, say), the variance would come out in \$2.

Standard Deviation

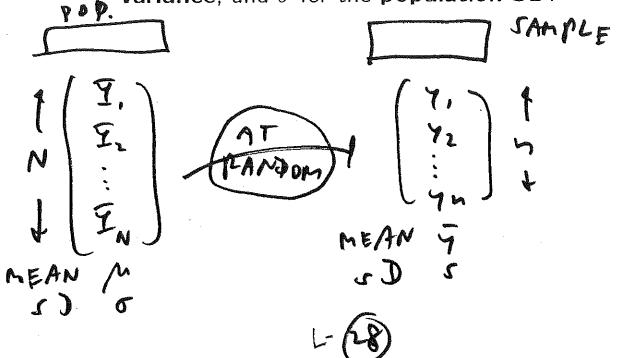
Both of these problems can be **solved** by taking the **square root** of the variance, which is called the **standard deviation** (**SD** for short), usually abbreviated s; for the **sample** this is

$$s = \sqrt{\frac{(y_1 - \bar{y})^2 + \dots + (y_n - \bar{y})^2}{n - 1}} = \sqrt{\frac{1}{n - 1} \sum_{i=1}^n (y_i - \bar{y})^2}.$$
 (4)

With the little **fake data set** the **sample SD** comes out $s = \sqrt{19} \doteq 4.4$, which seems **about right** as a **summary of spread** for this list of numbers.

Samples and populations. Suppose that the data set (y_1, \ldots, y_n) is a random sample of size n from a population of N values, and to keep from getting confused between the population and the sample let's call the population values (Y_1, \ldots, Y_N) ; we've already agreed to call the sample mean, variance and SD \bar{y} , s^2 and s, respectively.

With this notation it would be **natural** to call the **population mean, variance** and **SD** \bar{Y} , S^2 and S, respectively, but instead people typically use **Greek letters** and write μ for the **population mean**, σ^2 for the **population variance**, and σ for the **population SD**.



Graphical Interpretation of the SD

In this setup, to make things even more **mysterious**, people define the **population variance** and **SD** not by dividing by (N-1) but by N:

$$\sigma^2 = \frac{1}{N} \sum_{i=1}^{N} (Y_i - \mu)^2$$
 and $\sigma = \sqrt{\frac{1}{N} \sum_{i=1}^{N} (Y_i - \mu)^2}$. (5)

The **reason** for all of this **mystery** will be explained later when we talk about **sampling distributions**; for now it's enough to notice that when n is **large** it **hardly matters** in practice whether you divide by n or (n-1) in calculating the **sample SD** s.

Graphical interpretation of the SD. SDs are a pain to compute by hand or with a calculator, and it's easy to make mistakes when doing so, so it would be good to have a simple way to roughly approximate the SD of a list of numbers by looking at its histogram — this is provided by something called the Empirical Rule.

Empirical Rule: For almost any list of numbers, if you start at the mean and go one SD either way, you'll capture about $\frac{2}{3}$ of the data (the theoretical figure is 68%); if you start at the mean and go 2 SDs either way you'll get most of the data (the theoretical number to remember is 95%); and if you start at the mean and go 3 SDs either way you'll get almost all of the data (the theoretical figure is 99.7%).

Looking back at the **butterfly wing length** example, for instance (the **sorted data values** are on page 12, and the **histogram** is on page 13; recall that the **mean** is about 4 cm), you can see that if you guessed the SD was about 0.1 cm that would be too small (since starting at 4.0 and

1.4 Using the Normal Distribution Descriptively

going 0.1 either way, down to 3.9 and up to 4.1, would only get you half of the data), and if you guessed s=0.5 that would be too big, because the interval from (mean – 1SD) = (4.0 - 0.5) = 3.5 to (mean + 1SD) = (4.0 + 0.5) = 4.5 includes all of the data; a bit of trial and error should convince you that the SD is around 0.3 cm, and actually computing it yields s=0.29 cm.

Using the normal distribution descriptively. There's a distribution that's quite special in many ways (you've probably already met it) — it's the symmetric unimodal distribution called the normal (or Gaussian) distribution.

Actually there's **not just one normal distribution**, there are **infinitely many** of them: for any values \bar{y} and s you can imagine for the **mean** and **SD** (respectively) of a **single-variable data set** (thought of as a **sample** from a **population**), there's a **normal distribution** with that **mean** and **SD**.

Looking at the histogram of the butterfly wing length data (page 13), which is roughly symmetric and unimodal, you can imagine someone approximating it with a smooth curve drawn through or near the tops of the bars — the curve (or function) corresponding to the normal distribution with mean $\bar{y} = 4.0$ cm and SD s = 0.29 cm has the equation (with y as the values of the variable running along the horizontal axis)

$$f(y) = \frac{1}{s\sqrt{2\pi}} \exp\left[-\frac{(y-\bar{y})^2}{2s^2}\right]. \tag{6}$$

The Normal Curve

The idea behind using the normal distribution (also called the normal curve) descriptively is as follows.

Suppose you wanted to answer a question like "What percentage of the butterflies in this data set had wing lengths smaller than 3.56 cm?"

The **exact** way to answer this question is just to **count** how many butterflies satisfied this criterion (2, as it turns out: the ones with wing lengths of 3.3 and 3.5 cm) and see what **percentage** this is of the **total sample size**; here, the answer would be $\frac{2}{24} = 0.083 = 8.3\%$.

To use the **normal curve** to get an **approximate** answer to this same question, we can reason as follows.

All histograms express **frequency** information, but it turns out that there are **three different possible choices** of the **vertical** axis for histograms, and each choice expresses the idea of frequency in a **different** way.

The histogram back on page 13 was a <u>raw frequency</u> histogram — the <u>vertical axis</u> just plotted the <u>raw frequencies</u> (counts) of data values in each bar.

Another idea would be to divide the raw frequencies by the sample size, to produce relative frequencies, and plot them on the vertical scale instead — this would give you a relative frequency histogram, which would have the same shape as the raw frequency picture (technically speaking, in moving from raw to relative frequency the vertical axis has undergone a linear change of scale, and that has no effect on the shape of the basic distribution).

The Normal Curve (continued)

There's a **third way** to plot histograms: to draw the **vertical axis** on what's called the **density scale**, which is chosen so that

- (a) **relative frequency** is expressed by computing the **area** under the histogram or curve, and
- (b) the **total area** under the histogram or curve is therefore **100%** (or **1**).

When considered as an approximation to a histogram, the normal curve (as it turns out) is (by definition) drawn on the density scale, so it can be used to approximate relative frequencies (like the percentage of the data values less than 3.56) by working out the area to the left of 3.56 under the normal curve with the same mean and SD as the data.

Q: How do you work out the area under a normal curve?

A: Formally speaking, the calculus technique of integrating the function in equation (6) from $-\infty$ to 3.56 will give you the right answer, but it turns out that the Gaussian density function in (6) cannot be integrated in closed form (it has no anti-derivative), so the back-up technique is called numerical integration: you use numerical (rather than symbolic) methods to make a table in which the area under the normal curve to the left of some number z is computed for lots of different choices of z.

Q: But you said earlier that there isn't just one normal curve, there are infinitely many of them, one for each choice of the mean and SD, so with this approach wouldn't you have to create infinitely many tables?

The Normal Curve (continued)

A: Ah, yes, embarrassing but good question; to rescue this idea we have to appeal to a remarkable fact about the normal distribution:

Fact: Every normal curve (no matter what it's mean and SD is) satisfies the (theoretical version of the) Empirical rule not just approximately but exactly; in other words, if you start at the mean and go 1 SD either way, the area under any normal curve will be 68%; 2 SDs either way, 95%; and 3 SDs either way, 99.7%.

This means that it's enough to make a **table** for **only one normal curve** — by convention, it's called the **standard normal curve** — and then **relate** whatever normal curve you're interested in **back** to the **standard curve**.

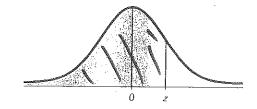
Q: What did people choose for the mean and SD of the standard normal curve?

Well, the **mean** could be anywhere from $-\infty$ to ∞ , and the **simplest number** between these **extremes** is probably $\mathbf{0}$, and the **SD** could be anywhere from 0 to ∞ (why can't an SD be negative?), and the **simplest number** in this range is probably $\mathbf{1}$, so the **standard normal curve** (by definition) has **mean 0** and **SD 1**.

The table on the next two pages (Table A-2 in T&T, also available on the inside back cover of the book) gives areas to the left of a bunch of places z under the standard normal curve; for example, to work out the area to the left of -1.27 you look in the row marked -1.2 and the column marked 0.07 (because ignoring the minus sign and putting the 1.2 and the 0.07 together you get 1.27), and the table says the area is 0.1020, which could also be expressed as 10.20% (in practice this would typically be rounded to 10.2% or 10%, because the normal curve is only being used as an approximation to the actual histogram).

The (Standard) Normal Table

POSITIVE z Scores



TABLE/A	2 (co	ntinued) Ci	umulative	Area from	the LEFT					
Z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
0.0	.5000	.5040	.5080	.5120	.5160	.5199	.5239	.5279	.5319	.5359
**:0.1	5398	.5438	.5478	.5517	5557 _E	.5596	5636	5675	.5714	.5753
0.2	.5793	.5832	.5871	.5910	.5948	.5987	.6026	.6064	.6103	.6141
/ 7/0.31 V	:6179	6217	6255	6293	.6331	:6368	6406	%°6443°÷	6480	6517
0.4	.6554	.6591	.6628	.6664	.6700	.6736	.6772	.6808	.6844	.6879
.O.5	6915	.6950	6985	.7019	7054	.7088	7123 =	景.7157。	.7190	.7224
0.6	.7257	.7291	.7324	.7357	.7389	.7422	.7454	.7486	.7517	7549
* 0.7 · ·	7580	7611	7642	7673	7704	7734	7764	.7794	.7823	7852
0.8	.7881	.7910	.7939	.7967	.7995	.8023	.8051	.8078	.8106	.8133
0.9	8159	8186	.8212	J 8238	. 8264	8289	v8315	:8340	.8365	.8389
1.0	.8413	.8438	.8461	.8485	.8508	.8531	.8554	.8577	.8599	8621
1.1	.8643	.8665	8686	8708	.8729	.8749"	8770	.8790	8810	8830
1.2	.8849	.8869	.8888	.8907	.8925	.8944	.8962	.8980	.8997	.9015
1.3	.9032	.9049	.9066	.9082	.9099	.9115	9131	.9147	9162	9177
1.4	.9192	.9207	.9222	.9236	.9251	.9265	.9279	.9292	.9306	9319
1.5	.9332	.9345	.9357	.9370	.9382	.9394	9406	9418	9429	9441
1.6	.9452	.9463	.9474	.9484	.9495	* .9505	.9515	.9525	.9535	.9545
1.7	.9554	.9564	.9573	.9582	.9591	9599	9608	.9616	9625	.9633
1.8	.9641	.9649	.9656	.9664	.9671	.9678	.9686	9693	.9699	.9706
1.9	9713	.9719	9726	.9732	.9738	.9744	9750	9756	.9761	.9767
2.0	.9772	.9778	.9783	.9788	.9793	.9798	.9803	.9808	.9812	.9817
2.1	.9821	9826	.9830	.9834	.9838	9842	. 9846	9850	.9854	.9857
2.2	.9861	.9864	.9868	.9871	.9875	.9878	.9881	.9884	.9887	.9890
2.3	9893	9896	9898	9901	9904	9906	9909	2.9911	.9913	9916
2.4	.9918	.9920	.9922	.9925	.9927	.9929	.9931	.9932	9934	.9936
2.5 · · ·	.9938	9940	9941	9943	9945	9946	9948	.9949 *	.9951	9952
2.6	.9953	.9955	.9956	.9957	.9959	.9960	.9961	9962	9963	.9964
3.2.7	9965	9966	.9967	.9968	.9969	9970	.0971	≨:.9972 :: ¹	9973	9974
2.8	.9974	.9975	.9976	.9977	.9977	.9978	.9979	.9979	.9980	.9981
29	5.9981	.9982	.9982	9983	.9984	9984	9985	9985	9986	9986
3.0	.9987	9987	9987	.9988	9988	.9989	9989	9989	.9990	.9990
3.1	29990		.9991	9991	9992	9992	9992	9992	.9993	.9993
3.2	.9993	.9993	9994	.9994	.9994	.9994	.9994	.9995	.9995	.9995
3.3	9995	(*** 9 995***	39995	9996	9996	39996	0996	9996	9996	9997
3.4	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9998
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NOTE: For values of z above 3.49, use 0.9999 for the area. *Use these common values that result from interpolation:

 z score
 Area

 1.645
 0.9500

 2,575
 0.9950

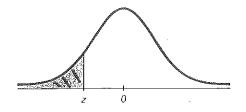
Common Cri	tical Values
Confidence	Critical
Level	Value
0.90	1.645
 0.95	1.96

2.575

0.99



The (Standard) Normal Table



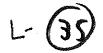
NEGATIVE z Scores

TABLE A-	2 . Stan	dard Norm	ial (z) Dis	tribution:	Cumulativ	e Area froi	m the LEF	Γ		
. Z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
-3.50									-	
and	1									
lower	.0001						1			
-3.4	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0002
3.3	:0005	-:0005	.0005	.0004	0004	.0004	.0004	0004		.0003
-3.2	.0007	.0007	.0006	.0006	.0006	.0006	.0006	.0005	.0005	.0005
3.13%	.0010	.0009	÷.0009	0009	.0008	8000	.0008	0008	.0007	= .0007
-3.0	.0013	.0013	.0013	.0012	.0012	.0011	.0011	.0011	.0010	.0010
-2.9 🐒	.0019	.0018	* .0018: ₁₋₁		0016	.0016	.0015	.0015		.0014
-2.8	.0026	.0025	.0024	.0023	.0023	.0022	.0021	.0021	.0020	.0019
-2.7	.0035	.0034	0033	0032	0031	.0030	.0029	0028	, .0027	0026
-2.6	.0047	.0045	.0044	.0043	.0041	.0040	.0039	.0038	.0037	.0036
-2.5	.0062	\$.0060	.0059	.0057	- 10055	.0054	.0052	.0051	* .0049	.0048
-24	.0082	.0080	.0078	.0075	.0073	.0071	.0069	.0068	.0066	.0064
-2.3	.0107: 5		0102	.0099	· [.0096.]	.0094	.0091	.0089	ै .0087	.0084
-2.2	.0139	.0136	.0132	.0129	.0125	.0122	.0119	.0116	.0113	.0110
	.0179	.0174	.0170	.0166	0162	0158	-0154	40150	.0146	.0143
-2.0	.0228	.0222	0217	.0212	.0207	.0202	.0197	.0192	.0188	.0183
-1.9	.0287	.0281		control to the property of the control of the contr	0262	.0256	0250	.0244	.0239	0233
-1.8	.0359	.0351	.0344	.0336	.0329	.0322	.0314	.0307	.0301	.0294
-1.7	-:0446	.0436	.0427	.0418	.0409	.0401		0384	.0375	.0367
-1.6	.0548	.0537	.0526	.0516	.0505	* .0495	.0485	.0475	.0465	.0455
-1.5	.0668		.0643	.0630	£0618 £	.0606	.0594	0582	0571	0559
-1.4	.0808	.0793	.0778	.0764	.0749	.0735	.0721	.0708	.0694	.0681
··· +1.3, 📉	0968	.0951	∴0934	.09183	300003	.0885	.0869	0853	∜.0838	0823
-1.2	.1151	.1131	.1112	.1093	.1075	.1056	.1038	.1020	.1003	.0985
· -1.13	1357	1335	.1314	.1292	11271	1251	1230	.12105.	1190,	.1170
-1.0	.1587	.1562	.1539	.1515	.1492	.1469	.1446	.1423	1401	.1379
÷÷0.9 ∗∜	:1841	1814	1788	1762	1796%	\$\$1711\frac{1}{2}	1685	1660	18.1635	-16110
-0.8	.2119	.2090	.2061	.2033	.2005	.1977	.1949	.1922	.1894	.1867
-0.7	:2420-s.	¥.2389	.2358	2327	第2296 元	2266.	.2236	2206	器,2177	.2148
-0.6	.2743	.2709	.2676	.2643	.2611	.2578	.2546	.2514	.2483	.2451
#÷-0,5 €	13085	£13050	C:3015	22981	2996	2912	2877	2843	2810	2776
-0.4	.3446	.3409	.3372	.3336	.3300	.3264	.3228	.3192	.3156	.3121
3 - 0.3	.3821	3783	3745 公	20707	20069	3632	3594	3557,74	3520	.3483
-0.2	.4207	.4168	.4129	.4090	.4052	.4013	.3974	.3936	.3897	.3859
÷ +0.1	.4602	4562	.4522		* (A143 ·	2.4404	. 4364	4325	/ ₂ .4286	.4247
-0.0	.5000	.4960	.4920	.4880	.4840	.4801	.4761	.4721	.4681	.4641
										

NOTE: For values of z below -3.49, use 0.0001 for the area.

^{*}Use these common values that result from interpolation:

z score	Area
-1.645	0.0500
2 575	0.0050



The Normal Curve (continued)

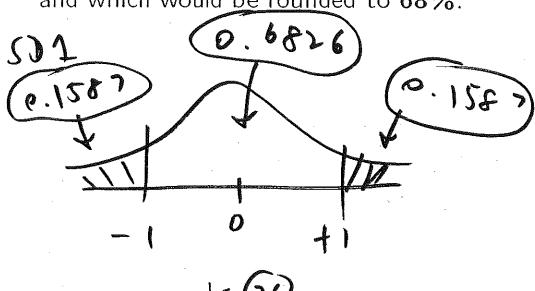
Q: The table only gives areas to the left of z under the standard normal curve; what if I want the middle area between -z and +z, or the area to the **right** of z?

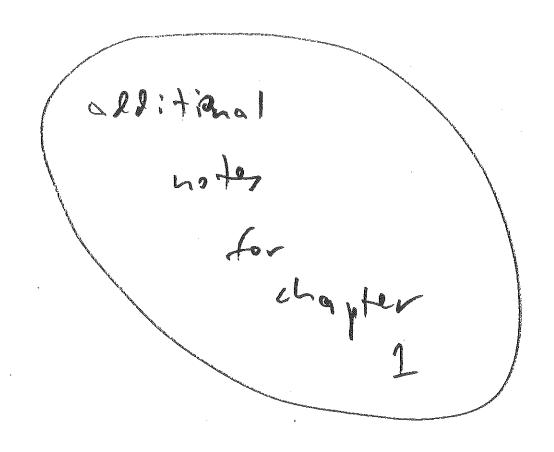
A: You can use two basic facts about the normal curve to work out any other kind of area you want:

- Since all normal curves are (by definition) drawn on the density scale, the total area under any normal curve is 100% (or 1), and
 - All normal curves are symmetric around their means.

For example, to work out the area between -1 and +1under the standard normal curve, this corresponds to going 1 SD either way from the mean (0), so we know from the theoretical figure in the Empirical Rule that the answer should be 68%.

Reasoning from the table, you can look up the area to the **left of -1** and get 0.1587; by symmetry the area to the right of +1 must also be 0.1587; and because the total area is 1 the area in the middle must be 1 - 2(0.1587) = 0.6826, which we might express as 68.26%and which would be rounded to 68%:





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1 Vu	High	High	A Code
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same conter, o. Ferst shape center (4.4)

1 n=24 4.4+...+3.9 = 3.96 cm "y bar

place of highest frequency X La Mean 2 balance yeint hardian of willenot hato sorted for smallest to 1979t - & place when o have = Ato below, I above 50% 50% h=3 [1] 50+ (2) mean 4 1 2 50xt (2) 1=4 bufferfly: 1-63 melian 2+6 = 4

Vtof gnut mode penentile with his 50 12 penetik 12/4/2

L (54)

(7,-7)+(7,-7)+...+(7,-7) 2 (4)+72+...+72) - (9+9+...+7) - エアン・ハイ = エアン・メ(大ニアン) - (ビィンノー(ビアン) = ロ ne data vivalize / now [n]. nominal: Ny lifter dichotomoys V S(1=4,0=N) (a) (quant) (discrete) J (dicho tomour)

(b) phosphate concentration I vou for each stream location n=60 2hont Continuous (n+i3) (c) tem (°C) et which chirpme fatte telem 100/min 24.8 9 1 vow for each night 27.2 1 = 44 75° F °F = 2°C + 32° (2 mont) (cont) 20°C = 18°F (interval) Kind of animal 25°C = 77°F frutte frutte frutte h=? vertesuate amin qual) (nominal) (not 1:ch) (5)

DS 2 | #2 °F = 2°C + 32° nultirizing a dataset by a constact her no effect on besie shope of hist:

Head 19

Kist:

Here

if wat 2

Here

if wat 2

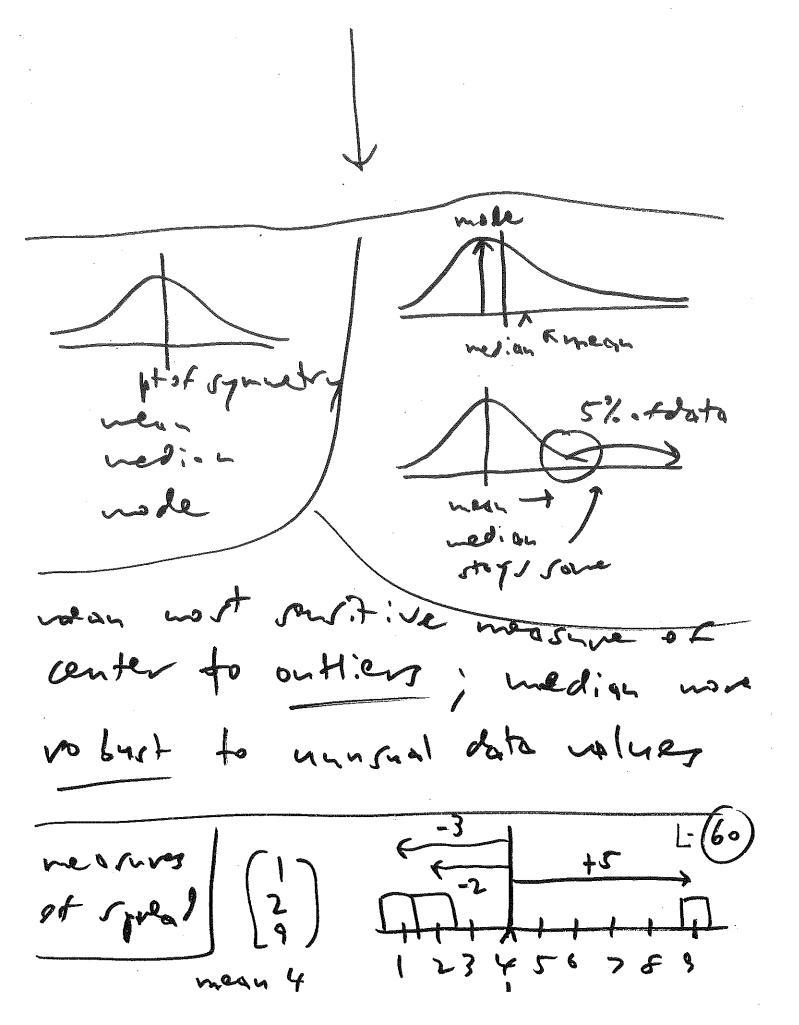
wear jets waltiplied by 2 & speed also los ul adding a constant to all Joto valuer (II) +7 (men)

just whifts

hirt to lefts (vagetive courtout) 12

or tight (vagetive courtout) 12 or tight (possitive constant) by the ansunt of the bustont, ulich again has no effect on shape of hist; pe men jos my (quant A) ~ 1. mg (bust 0) by the unsunt of the constant, but spread is unchonjed. (8)

So if I wearne temp it F & you use e our hist will have exactly the squal shape 3.3 3.4 3.5 7.6 7.7 3.8 3.9 4.0 4.1 4.2 4.3 4.4 4.5 Mlues h=24, a better



typical length of among + $\begin{pmatrix} 1-4 \\ 2-4 \end{pmatrix} = \begin{pmatrix} -3 \\ -2 \\ +5 \end{pmatrix} \begin{pmatrix} to set \\ till of \\ terial to see \\ for weak \end{pmatrix}$ i sewal [] 7: -7] (AAD) $2\left(\frac{3}{2}\right)$ mean (3.3) average assolute deciation (MA) (kr 1-1 55 nars: ye ha 1294 mean 84 L- (6)

new (9) the 5° of Ph.; Somple = \(\lambda \) \(\lam sa-ple variouse L [(():-7)] = 5 = paphical interpretation with nice, et 1): [Empirical Rule: start st near of almost any alta set, po 1 (2)[3] SD either may: you with wra ally capture

about 1/3 (most) (almost) glust no møter det hist. My Minde by somple 50 55 Y. Free men K Lepes of treedon for wearing 1505

givet % . Flata ~1 62/2w 7.56 cm? 3.5640 2 of 24 who we . 3.5% so this % is 24 (exoch) A: (exect) Lirt ir 23.56? Loh (syproximate) standed hornal 5- (T. E(2: -7), 50

no mad card + 68% 5- Heafly Tofts raw units axis (4) .561 4.0cm -1.25-1 - 4.0 g/m = 2 score to Stocker 54) (7= + 2.5) conerty

